Time Series and Extreme Value theory WI4230 Final Exam - June 2016

Please provide your student card on the table, ready for inspection.

Mobiles, tablets and similar objects must be switched off. The exam is invalidated if you cheat.

Please write with a pen. Please write your name, surname and student number on all your papers.

The exam is made of two parts. You have in total 3 hours.

If you sat for the Time Series' Partial in April, and you like your grade, you can skip the Time Series Part. In that case, you then have 1.5 hours

Good Luck!

Time Series Part

Please answer 4 out of the following 5 questions. They all have the same weight. You can choose the ones you prefer.

- 1. Consider the following points.
 - Give the formal definitions of strict stationarity and weak stationarity.
 - Provide at least one example of time series for each definition.
 - List at least three possible causes of non-stationarity.
- 2. Consider the process

$$x_t = \beta_0 + \beta_1 t + w_t,$$

where β_0 and β_1 are fixed constants, and w_t is a white noise with zero mean and variance σ_w^2 .

- Is x_t strictly and/or weakly stationary? Prove or disprove it.
- Prove that the first difference series $\nabla x_t = x_t x_{t-1}$ is weakly stationary.
- Repeat the previous point, when w_t is replaced by a general stationary process, say y_t , with mean function μ_y and autocovariance function $\gamma_y(h)$.
- **3.** Define the process

$$x_t = \phi x_{t-1} + w_t,$$

where $|\phi| > 1$ and $w_t \sim WN(0, \sigma_w^2)$.

- Compute the autocovariance function of x_t .
- Verify that x_t is not a causal process.
- Define a process y_t , equivalent to x_t , which is causal.
- 4. Using the operator representation, we define a process

$$(1 - 0.4B - 0.45B^2)x_t = (1 + B + 0.25B^2)w_t$$

where w_t is the usual white noise $w_t \sim WN(0, \sigma_w^2)$.

- What type of model is this? Is the above representation redundant? Motivate your answer.
- Is the process causal? Verify and, in case, give the corresponding representation.
- Is the process invertible? Verify and, in case, give the corresponding representation.
- **5.** Let $w_t \sim WN(0, \sigma_w^2)$ be a white noise.

• Compute the spectrum (spectral density) of a TS x_t , where x_t is a MA(1) of the form

$$x_t = w_t + \frac{1}{2}w_{t-1}.$$

• Compute the spectrum of a TS x_t , where x_t is an AR(2) of the form

$$x_t = x_{t-1} - \frac{9}{10}x_{t-2} + w_t.$$

EVT Part

1. (6 points) Let V be a uniform random variable on [0,1] and

$$X = \frac{\frac{1}{V}}{\log \frac{1}{V}}.$$

Show that $X \in D(G_1)$, that is X being in a max domain of attraction with extreme value index 1.

- **2.** Let $G_0(x) = \exp(-\exp(-x)), x \in R$.
 - (5 points) Let $E_{1,n} \leq E_{2,n} \leq \ldots \leq E_{n,n}$ be the order statistics of a random sample from Exp(1). The distribution function of Exp(1) is given by $F(x) = 1 - \exp(-x)$. Find $a_n > 0$ and $b_n \in R$ such that

$$\lim_{n \to \infty} \Pr\left(\frac{E_{n,n} - b_n}{a_n} \le x\right) = G_0(x).$$

• (5 points) Let $\{X_i, i=1,\ldots,n\}$ be a random sample with distribution function F(x)= $1 - x^{-\frac{1}{\gamma}}$, x > 0 and $\gamma > 0$. Prove that if $k = k(n) \to \infty$ and k < n, then as $n \to \infty$,

$$\frac{\log X_{n,n} - \log X_{n-k,n}}{\log k} \xrightarrow{P} \gamma$$

where $X_{1,n} \leq X_{2,n} \leq \ldots \leq X_{n,n}$ are the order statistics. **Hint:** You might use that $\{E_{n-i,k} - E_{n-k,n}\}_{i=0}^{k-1} = d \{E_{k-i,k}^*\}_{i=0}^{k-1}$, where $\{E_1^*, \ldots, E_k^*\}$ is a random sample from Exp(1).

3. Based on a random sample X_1, \ldots, X_n from an unknown distribution F, one would like to build up an estimation procedure for the following probability,

$$\Pr(X_1 > x_0),$$

where x_0 is a given large value. Let U be the tail quantile function: $\Pr(X_1 > U(t)) = \frac{1}{t}$, for $t \ge 1$. Assume the max domain of attraction condition on F, that is there exists a positive function a(t)and $\gamma \in R$ such that for x > 0,

$$\lim_{t \to \infty} t(1 - F(a(t)x + U(t))) = (1 + \gamma x)^{-\frac{1}{\gamma}}$$
(1)

• (5 points) Show that (1) is equivalent to

$$\lim_{s \to x^*} \Pr\left(\frac{X_1 - s}{f(s)} > x | X_1 > s\right) = (1 + \gamma x)^{-\frac{1}{\gamma}},\tag{2}$$

where x^* is the right endpoint of F and f(t) a positive function.

• (4 points) Justify the following approximation

$$\Pr(X > x_0) \approx p \left(1 + \gamma \frac{x_0 - U(1/p)}{f(U(1/p))} \right)^{-\frac{1}{\gamma}}.$$

What conditions on p are needed for this approximation?

- (5 points) For a chosen p, how do you estimate γ , U(1/p) and f(U(1/p))? Describe the estimation procedure.
- **4.** Let (X_1, X_2) be a bivariate random vector with a continuous distribution function F. Let U_1 and U_2 the tail quantile functions: $\Pr(X_i > U_i(t)) = \frac{1}{t}$, for $t \ge 1$ and i = 1, 2. Suppose there exists a non-degenerate distribution G such that for $(x, y) \in (0, \infty] \times (0, \infty]$,

$$\lim_{n \to \infty} F^n(U_1(nx), U_2(ny)) = G(x, y).$$

- (3 points) Compute $G(x, \infty)$, for x > 0.
- (3 points) Show that for any (x, y) for which 0 < G(x, y) < 1,

$$\lim_{n \to \infty} n (1 - F(U_1(nx), U_2(ny))) = -\log G(x, y).$$

• (4 points) Define $L(x,y) = -\log G(\frac{1}{x}, \frac{1}{y}), \ x,y > 0.$ Show that for positive x and y, $\max(x,y) \le L(x,y) \le x + y.$